

WEIGHTED CLOSED-FORM ESTIMATORS FOR BLIND SOURCE SEPARATION

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Abstract

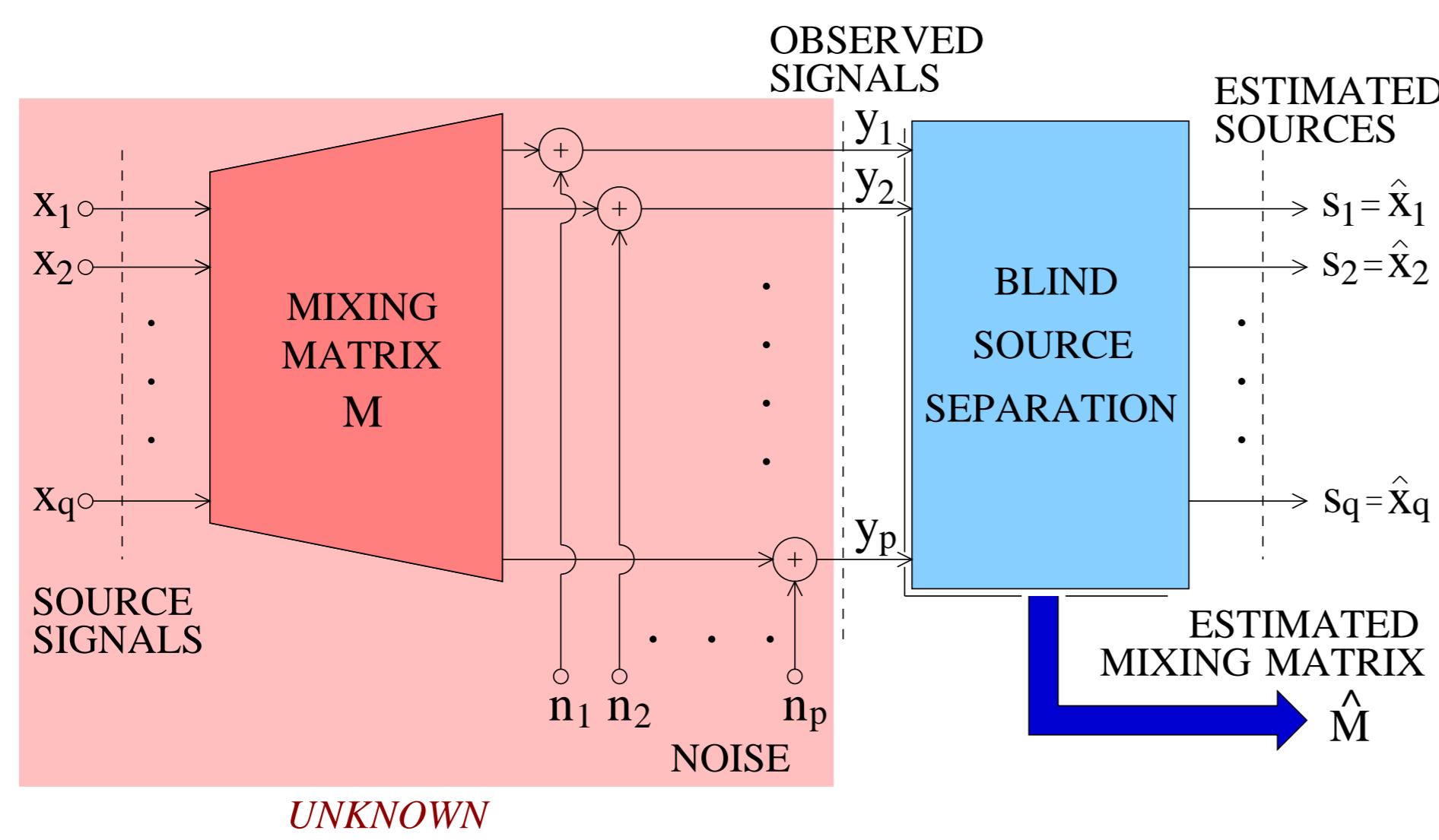
This paper investigates a novel closed-form estimation class, so-called weighted estimator (WE), for blind source separation in the basic two-signal problem. Proper combination of previously proposed estimators yields consistent estimates of the separation parameters under general conditions. In the real-mixture case, we determine analytic expressions for the WE asymptotic (large-sample) variance and the source-dependent weight value of the most efficient estimator in the class. By means of the bicomplex-number formalism, the WE is extended to the complex-mixture scenario, for which Cramér-Rao bounds are also derived. Simulations compare the WE with other methods, demonstrating its potential.

PROBLEM

- **Blind source separation (BSS)** of instantaneous linear mixtures:

$$\mathbf{y} = \mathbf{M}\mathbf{x}, \quad \begin{cases} \mathbf{y} \in \mathbb{C}^p : & \text{sensor array output} \\ \mathbf{x} \in \mathbb{C}^q : & \text{independent source signals} \\ \mathbf{M} \in \mathbb{C}^{p \times q} : & \text{mixing matrix.} \end{cases}$$

- **Objective:** from only knowledge of \mathbf{y} → estimate \mathbf{x} and \mathbf{M} .



REAL-MIXTURE CASE

Fourth-Order Weighted Estimator

- **Idea:** complex linear combinations (*centroids*) of whitened-sensor HOS → explicit expressions for estimation of θ .
- **EML:** $\xi_4 = (\kappa_{40}^z + \kappa_{04}^z - 6\kappa_{22}^z) + j4(\kappa_{31}^z - \kappa_{13}^z) = \gamma e^{j4\theta} \Rightarrow \hat{\theta}_{\text{EML}} = \frac{1}{4}\angle(\hat{\gamma}\hat{\xi}_4)$
- **AEML:** $\xi_2 = (\kappa_{40}^z - \kappa_{04}^z) + j2(\kappa_{31}^z + \kappa_{13}^z) = \eta e^{j2\theta} \Rightarrow \hat{\theta}_{\text{AEML}} = \frac{1}{2}\angle\xi_2$
- **AML, MaSSFOC:** combination of EML and AEML.
- **WE:** $\xi_{\text{WE}} = w\gamma\xi_4 + (1-w)\xi_2^2 = (w\gamma^2 + (1-w)\eta^2)e^{j4\theta} \Rightarrow \hat{\theta}_{\text{WE}} = \frac{1}{4}\angle\xi_{\text{WE}}$

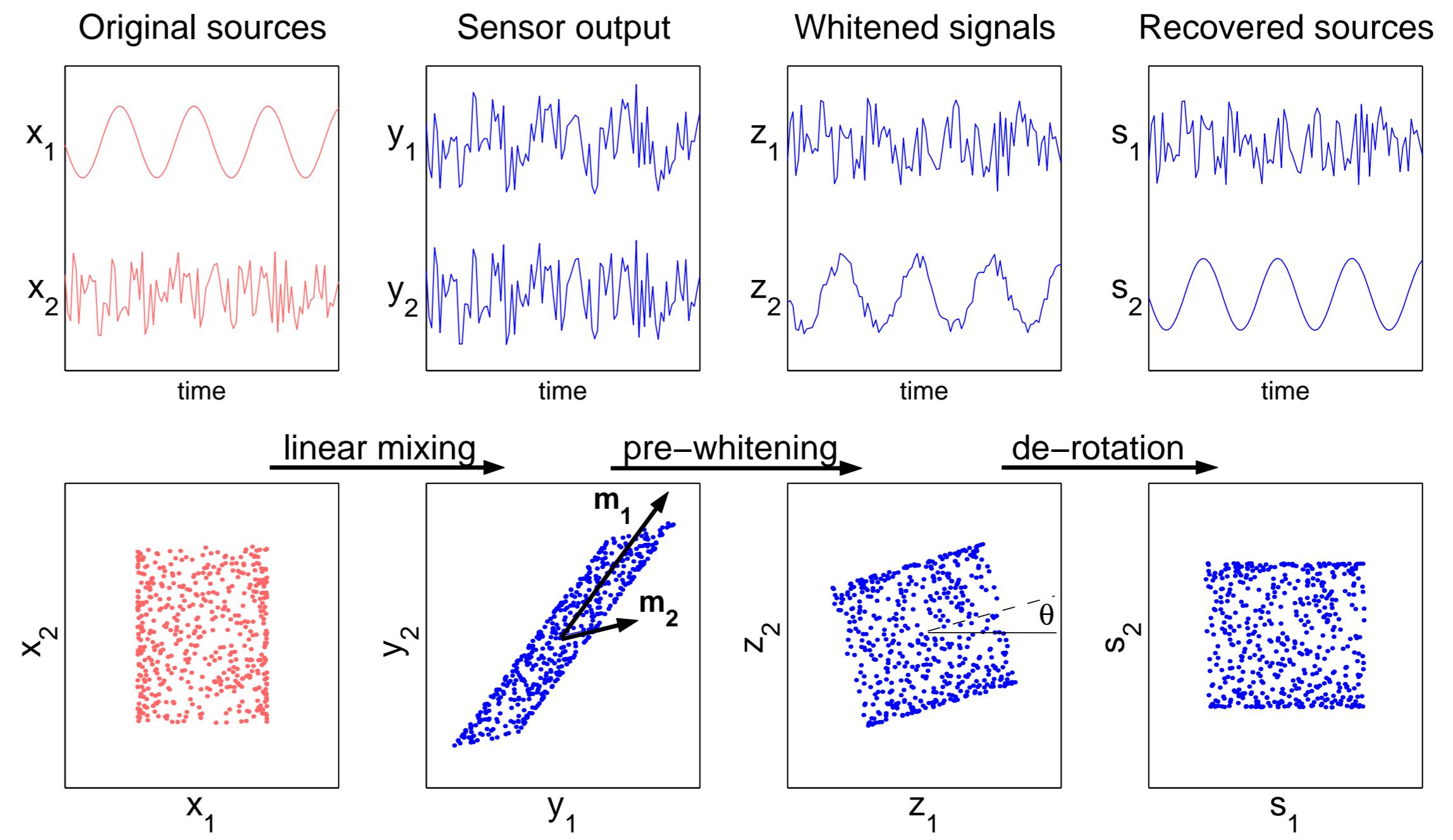
– Consistent for any source distribution, $0 < w < 1$.

– Optimal choice of w ?

- After 2nd-order spatial whitening: $\mathbf{z} = \mathbf{Q}\mathbf{x}$, $\mathbf{Q} \in \mathbb{C}^{q \times q}$ unitary.

- Fundamental two-signal case ($p = q = 2$):

$$\mathbf{Q} = \begin{bmatrix} \cos\theta & -e^{-j\alpha}\sin\theta \\ e^{j\alpha}\sin\theta & \cos\theta \end{bmatrix} \Rightarrow \text{estimation of } \theta, \alpha \in \mathbb{R}.$$



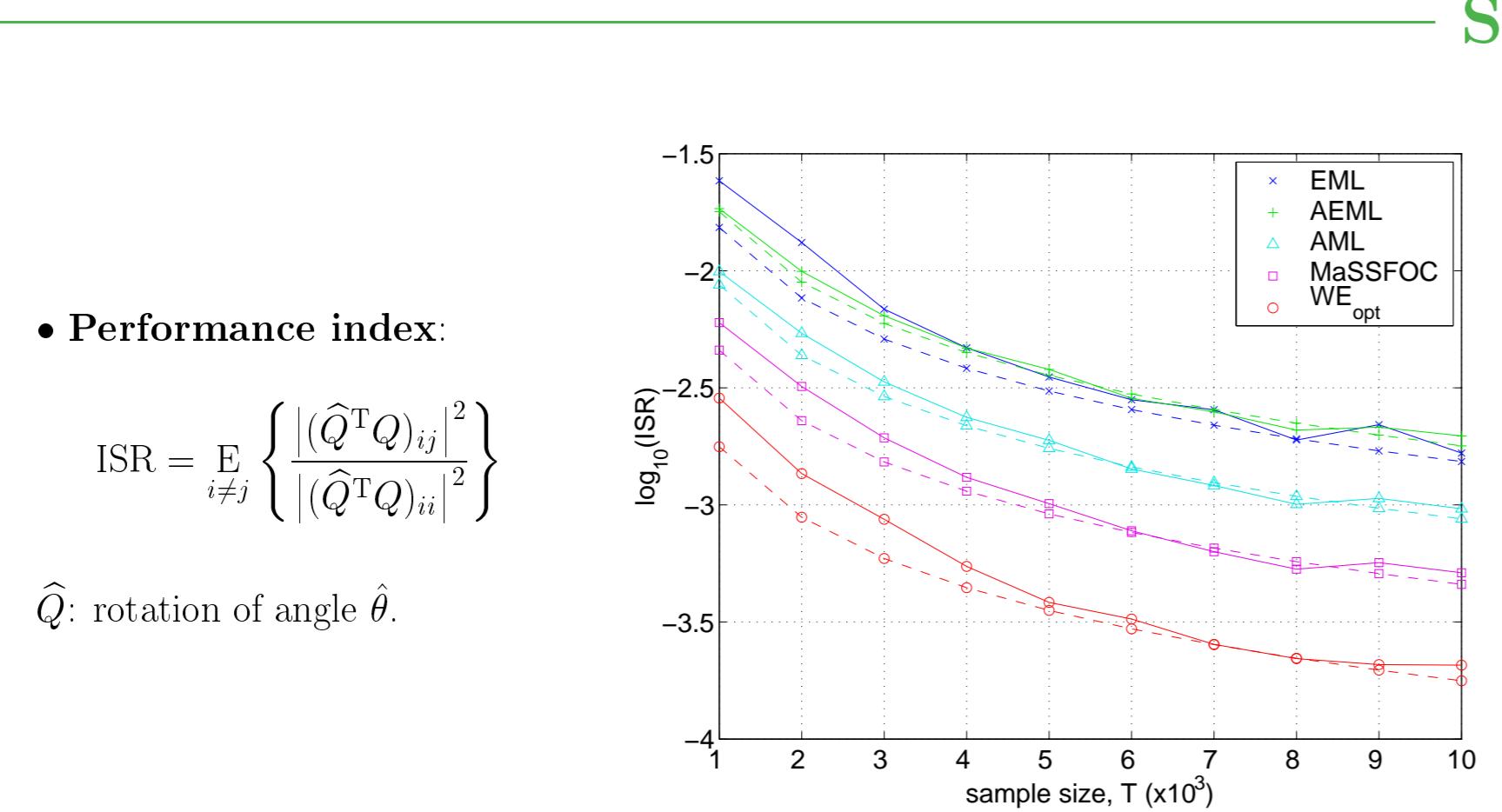
Optimal Finite-Sample Performance

- **Asymptotic (large-sample) variance of WE** (T samples):

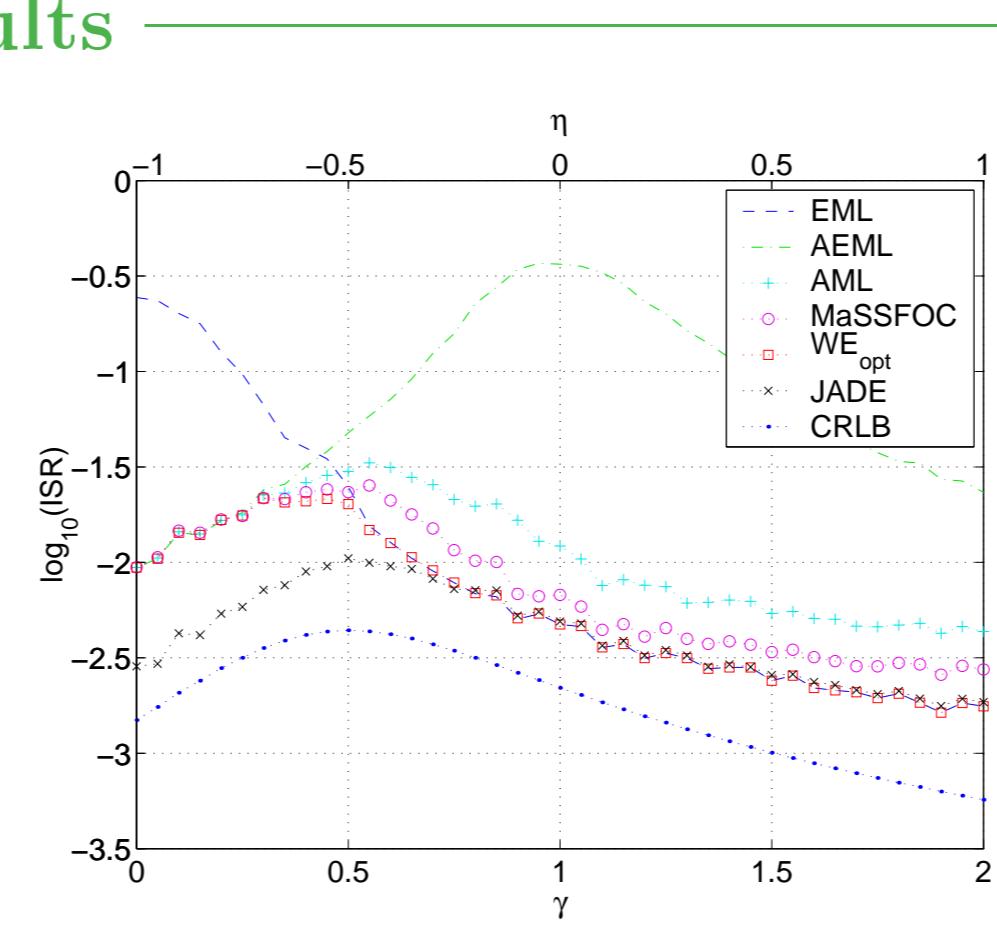
$$\sigma_{\theta_{\text{WE}}}^2 = \frac{\mathbb{E}\left\{ [w\gamma(x_1^3x_2 - x_1x_2^3) + (1-w)\eta(x_1^3x_2 + x_1x_2^3)]^2 \right\}}{T[w\gamma^2 + (1-w)\eta^2]^2}.$$

- **Minimum-variance WE** (if $|\kappa_{40}^x| \neq |\kappa_{04}^x|$):

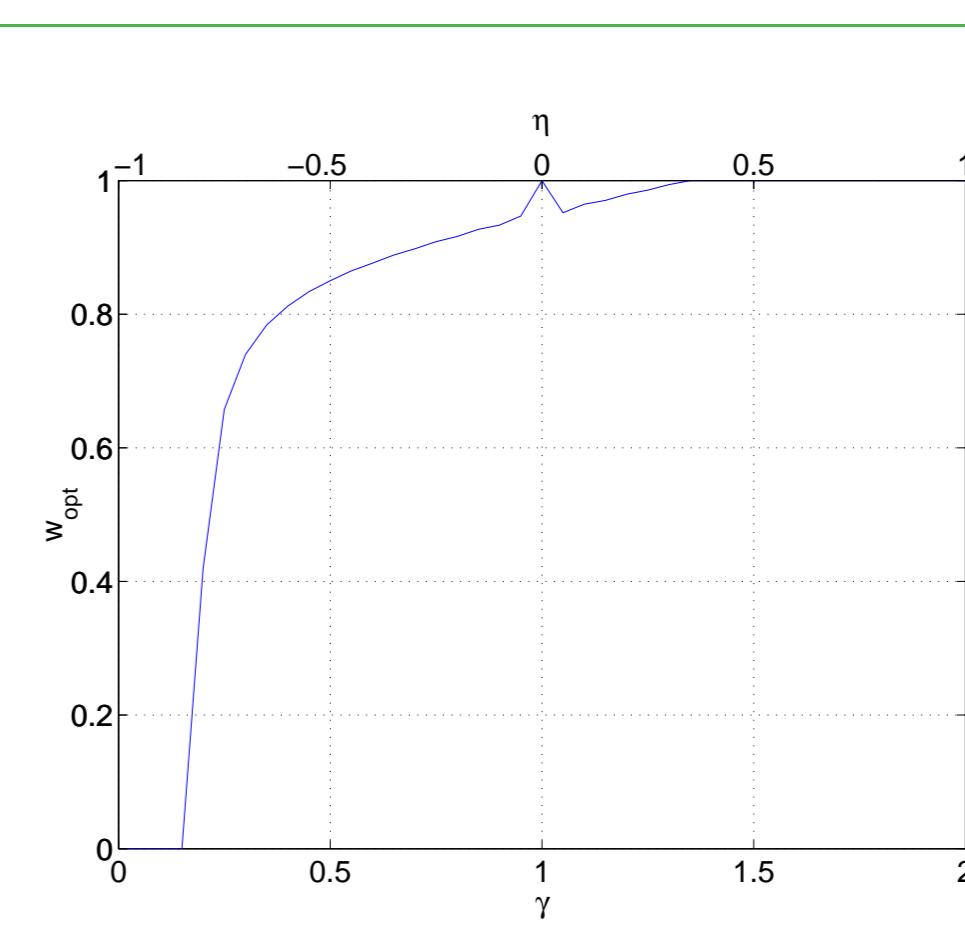
$$w_{\text{opt}} = \frac{1}{2} + \frac{\mu_{40}^x\mu_{04}^x[(\kappa_{40}^x)^2 - (\kappa_{04}^x)^2] + \kappa_{40}^x\kappa_{04}^x(\mu_{60}^x - \mu_{06}^x)}{2[(\kappa_{40}^x)^2\mu_{06}^x - (\kappa_{04}^x)^2\mu_{60}^x]}.$$



ISR vs. sample size. Uniform-Rayleigh sources, $\theta = 15^\circ$, ν independent Monte Carlo runs, with $\nu T = 5 \times 10^6$. Solid lines: average empirical values. Dashed lines: asymptotic variances.



ISR vs. sks γ and skd η . GGD sources, $\kappa_{04}^x = 0.5$, $\theta = 15^\circ$, $T = 5 \times 10^3$ samples, 10^3 Monte Carlo iterations.



Optimal value of the WE weight parameter.

COMPLEX-MIXTURE CASE

Bicomplex Numbers

- Unitary matrix $U = \begin{bmatrix} a & -b^* \\ b & a^* \end{bmatrix}$, $a, b \in \mathbb{C} \rightarrow$ bicomplex number $\bar{x} = a + jb$, $j^2 = 1$, $j \neq j$.
- Product: $\bar{x}_1 = a_1 + jb_1$, $\bar{x}_2 = a_2 + jb_2 \Rightarrow \bar{x}_1\bar{x}_2 = (a_1a_2 - b_1^*b_2) + j(b_1a_2 + a_1^*b_2)$.
- Unitary matrices under matrix product and bicomplex numbers under above product operation: isomorphic.
- Matrix $Q \rightarrow$ bicomplex exponential: $e_j^{\theta} = \cos\theta + je^{j\alpha}\sin\theta$.

Fourth-Order Weighted Estimator

- **Bicomplex centroids.**
- **CCEML:** $\bar{\xi}_4 = (\kappa_{40}^z + \kappa_{04}^z - 6\kappa_{22}^z) + j4(\kappa_{31}^z - \kappa_{13}^z) = \gamma e^{j4\theta}$.
- **CAEML:** $\bar{\xi}_2 = (\kappa_{40}^z - \kappa_{04}^z) + j2(\kappa_{31}^z + \kappa_{13}^z) = \eta e^{j2\theta}$.
- **CWE:** $\bar{\xi}_{\text{CWE}} = w\gamma\bar{\xi}_4 + (1-w)\bar{\xi}_2^2 = (w\gamma^2 + (1-w)\eta^2)e^{j4\theta}$.

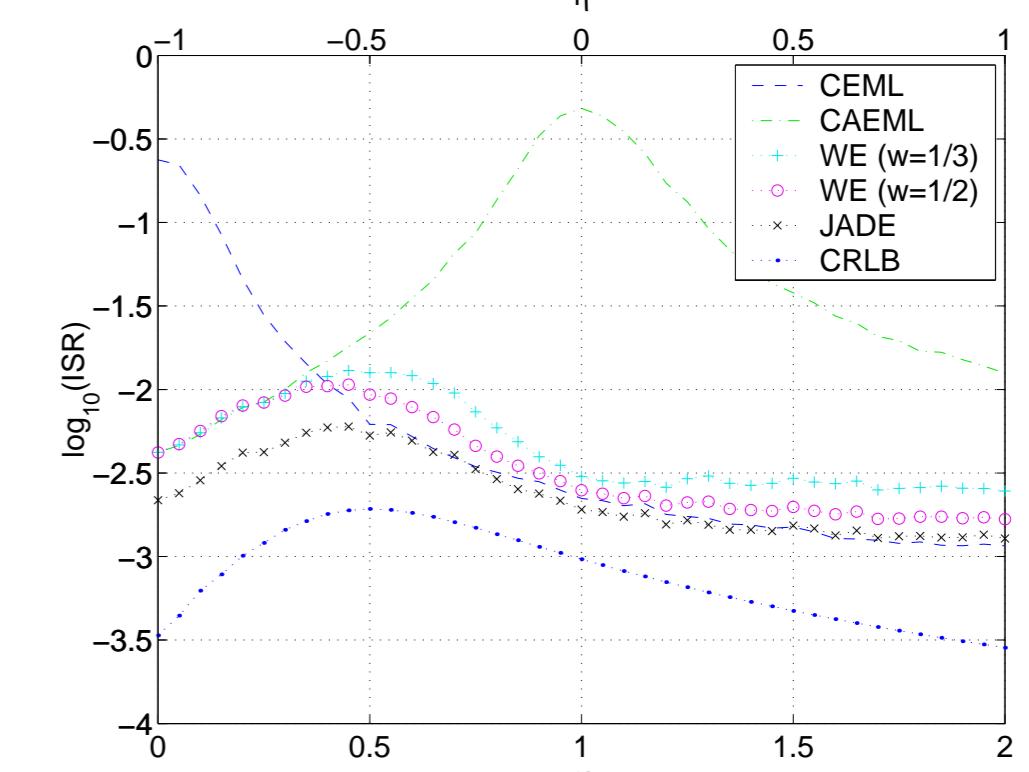
CONCLUSIONS AND OUTLOOK

- Linear combination of 4th-order centroids for closed-form BSS → consistent estimates of separation parameter for any source distribution ⇒ weighted estimator (WE).
- Prior knowledge of source statistics → WE with optimal finite-sample performance (minimum asymptotic variance).
- Bicomplex numbers → extension to complex mixtures. CRLB derived for circular sources. WE performance follows closely CRLB trend.
- **Further work:** – w_{opt} as a function of array-output statistics
 - behaviour in additive noise and impulsive interference
 - asymptotic performance analysis of complex WE

Cramér-Rao Lower Bounds

- Circularly distributed source signals, T samples: $\text{FIM}_{(\theta, \alpha)} = T \begin{bmatrix} I & 0 \\ 0 & \frac{1}{4}I \sin^2 2\theta \end{bmatrix}$, $I = I_1 + I_2 - 4$, $I_k = \frac{1}{2} \int \int \frac{1}{p_k} \left[\left(\frac{\partial p_k}{\partial u} \right)^2 + \left(\frac{\partial p_k}{\partial v} \right)^2 \right] du dv$, $p_k(u, v)$: pdf of the k th source signal $x_k = u_k + jv_k$, $u_k, v_k \in \mathbb{R}$, $k = 1, 2$.
- CRLBs of θ and α decoupled $\Rightarrow \begin{cases} \text{CRLB}_{\theta} = (TI)^{-1} \\ \text{CRLB}_{\alpha} = 4(TI \sin^2 2\theta)^{-1} \end{cases}$.
- Two Gaussian sources → $\text{FIM} = 0$.

Simulation Results



ISR vs. sks γ and skd η . CGGD sources, $\kappa_{04}^x = 0.5$, $\theta = 15^\circ$, $\alpha = 65^\circ$, $T = 5 \times 10^3$ samples, 10^3 independent Monte Carlo iterations.